



SYMPOSIUM

Student Journal of Science & Math

Volume 4, Issue 1

CAL POLY

SAN LUIS OBISPO
College of Science and Math

© *Copyright 2017*

*California Polytechnic
State University
University Graphic Systems
San Luis Obispo, California*



*To the students of Cal Poly who ask
the hard questions and rise to the
challenge of “learn by doing”*

Executive Editors

Jenna Vacca, *Editor-In-Chief*
Olivia Millay, *Managing Editor*
Clarisse Wangeline, *Managing Editor*
Tyler Scott, *Managing Editor*
Andrew Fulton, *Design Director*
Naba Maryam Ahmed, *Recruitment Editor*
Malamatenia Wilson, *Head Copy Editor*
Elise Barsch, *Copy Editor*
Bryce Aston, *Copy Editor*

Reviewers

Avity Norman
Leah Torres
Aiysha Mahmood
Camille Vernon

Advisor

Dr. Christopher Kitts

Table of Contents

| | |
|---------------------------------------|----|
| A Note from the Editor-in-Chief | 11 |
|---------------------------------------|----|

I. Research Articles

| | |
|--|----|
| Socioeconomic Status, Air Quality and Geographic Variation in Emergency Room Visits for Acute Bronchitis on the California Central Coast <i>by Sean Lang-Brown</i> | 15 |
|--|----|

II. Interdisciplinary Research

| | |
|--|----|
| Radiation Exposure During Space Travel: Using Radioisotopes for a Comparative Study of Human Feces and Urine as Integrated Shield Components <i>by Noah Falck</i> | 37 |
|--|----|

III. Special Feature

| | |
|------------------------------------|----|
| Letter from Dean Phil Bailey | 48 |
| End Note | 52 |

A Note from the Editor-in-Chief

Dear Readers of *Symposium*,

I am pleased to introduce the fourth installment of *Symposium: Student Journal of Science and Math*. The past year has been one of tremendous growth and change both within and around the publication. Allow me to bring you up to date.

For the first time ever, *Symposium* published two editions in one year, the first of which was the finalized 2016 issue. Although this was unconventional, the process proved that it is possible for future editions of the journal to be completed on a bi-annual basis.

This year also marked the expansion of the types of articles accepted to the journal. With the addition of two new sections – interdisciplinary research and abstracts – we seek to fulfill founder Lauren Young’s mission to spark scientific discussion through the inclusion of a wider variety of science-related student research. These additions will also enable us to keep up with the growing number of research opportunities occurring at Cal Poly.

We are also pleased to announce our first special feature article, a letter written by Dr. Phil Bailey, Dean of the College of Science and Math. Over the years, he has provided us with support, insight, and valuable feedback in planning for the future of *Symposium*, and we are truly grateful for the 48 years of service he has dedicated to our college. His effort in securing a \$110 million donation to the College of Science and Mathematics will open doors for the future generation of student researchers.

Additionally, this year marked the passing of the torch from our founding advisor Dr. Keeling to Dr. Kitts of the Biology department. We would like to thank Dr. Keeling for her enthusiasm and support of the journal since its inception. The staff of *Symposium* also wishes to offer a heartfelt welcome to Dr. Kitts. We know that *Symposium* will continue to find success under your guidance.

As always, thank you to Cal Poly for its funding via the Instructional Related Activities (IRA) program. Also, to Jeanine Scaramozzino, Dana Ospina, and Wilson Shao, thank you for your service to *Symposium* and your dedication to helping us publish a quality and accessible journal.

Now, without further ado, I formally present to the Cal Poly student community *Symposium: Student Journal of Science and Math*.

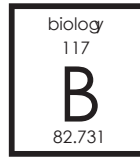
Sincerely yours,



Jenna Vacca

Editor-In-Chief, 2016-2017

I.
Research
Articles



Socioeconomic Status, Air Quality and Geographic Variation in Emergency Room Visits for Acute Bronchitis on the California Central Coast

By Sean Lang-Brown

Abstract

Importance

Analysis of geospatial variation in acute bronchitis due to socioeconomic and environmental factors can allow the efficient delivery of resources to populations most at risk.

Objective

We sought to determine if small-scale variation in socioeconomic factors and emergency room (ER) visits for acute bronchitis are associated with small cities or rural communities. We also modeled the effects of air quality on daily rates of ER visits for acute bronchitis in the context of socioeconomic factors to investigate modifying relationships.

Design, Setting, and Participants

We examined ER visits for acute bronchitis in San Luis Obispo and Santa Barbara counties from 2009 through 2012. The study area included 49 ZIP codes with a total population of 765,836 residents.

Exposures

Socioeconomic exposures included ZIP code level socioeconomic indicators collected for the 2010 American Community Survey. Environmental exposures included PM10, PM2.5, Ozone and temperature.

Main Outcomes and Measures

The rate of ER visits was calculated for each ZIP code. Spatial clustering (hotspots) of ER visits for acute bronchitis was examined using the local Getis-Ord G_i^* statistic. Differences between the distribution of socioeconomic variables across ER visit rate quintiles was assessed using the nonparametric Kruskal-Wallis test. Four Generalized Linear Mixed Models (GLMMs) were used to examine the association between lagged air quality, socioeconomic status and daily rates of ER visits for acute bronchitis in each ZIP code.

Results

5,620 ER visits for acute bronchitis were reported during the study period. The four-year rate of ER visits was between 2 and 17 visits per 1,000 residents for all ZIP codes. Two hotspots of ER visits were observed around the communities of Templeton and Lompoc, CA. Significant differences in home value and rent were observed across ER visit rate quintiles ($p = .003$ and $p < .001$, respectively). PM10 was found to be a significant predictor of daily ER visits in a GLMM including only environmental exposures. No exposures were found to be significant in a GLMM with both environmental and socioeconomic exposures. No clear evidence of socioeconomic factors modifying the effect of air quality on ER visits for acute bronchitis was found.

Conclusions

We found clear evidence of significant variation in ER visit rates for acute bronchitis at a small geographic scale in rural counties with small to medium size cities. Variation in ER visit rates across ZIP codes was associated with significant differences in socioeconomic factors including home value and rent.

Introduction

Socioeconomic status has been linked to the prevalence of respiratory infections on a small geographic scale. Recent work by Beck et al. has discovered high variation in the rate of pediatric emergency room visits for acute bronchitis and pneumonia across census tracts in one urban county near Cincinnati, Ohio, correlated with differences in socioeconomic status indicators.¹ The methods of Beck et al. have not been replicated in a rural area, and it is not reasonable to generalize these associations to communities that have different economic conditions from major cities, such as small cities, suburban and rural areas.

Air quality has been associated with respiratory infections and mortality in urban and rural communities.²⁻¹⁵ Daily air quality has been shown to predict emergency room visits days to weeks later for conditions including respiratory infections such as pneumonia and acute bronchitis, and air quality has been shown to affect mortality rates up to two months later.¹⁵⁻²¹

Evidence exists that socioeconomic factors are related to air quality effects on respiratory illness. Socioeconomic status has been found to modify the effect of air quality on respiratory mortality.²²⁻²⁴ However, little work has been done to observe socioeconomic effects in the context of daily emergency room visits for respiratory infections, especially in rural areas. It is unknown if socioeconomic factors modify the effect of air quality on respiratory infections (slope), or change the background rate of disease (intercept). To increase our understanding of the effects

of socioeconomic factors on respiratory infections, we performed two analyses in a rural area with small cities. We sought to determine if evidence of relationships between local variation in socioeconomic factors and emergency room visits for acute bronchitis can be found in small cities or rural communities. We also investigated the effects of air quality on emergency room visits for acute bronchitis in the context of local variation in socioeconomic factors to describe evidence of a modifying relationship. The California Polytechnic State University Institutional Review Board approved this study.

Methods

Study Location and Population

We examined the rates of emergency room visits for acute bronchitis in San Luis Obispo and Santa Barbara counties from 2009 through 2012. These counties are located on California's central coast. Land use is primarily agricultural with small towns to medium cities located along US Highway 101 and California Highway 1. We chose ZIP codes as the geographic unit of analysis to explore small-scale spatial variation in emergency room visit rates because the low population density in rural areas could confound finer spatial analysis. The study area included 49 ZIP codes with a total population of 765,836 residents after excluding three ZIP codes with less than 75 residents. Demographic characteristics of the study population can be seen in **Table 1**.

Emergency Room Visit Data

Acute bronchitis emergency room visits were defined using the International Classification of Diseases, Ninth Revision (ICD-9) code 4660.²⁵ Emergency room visit data was acquired from the California Office of State Health Planning and Development (OSHPD).²⁶ The dataset included emergency room visits with a primary diagnosis of acute bronchitis for residents of San Luis Obispo and Santa Barbara counties. The dataset was limited to date of emergency room visit, patient ZIP code of residence, age group, gender, secondary diagnosis and ethnicity (Hispanic or non-Hispanic) to protect patient privacy.

Environmental and Socioeconomic Data

Population size and socioeconomic data collected for the 2010 American Community Survey was downloaded from the US Census Bureau FactFinder Website.²⁷ Hourly air quality data was downloaded from the US EPA Air Quality System Data Mart.²⁸ Temperature data was downloaded from the California Agricultural Resources Board Meteorology Data Query Tool.²⁹ Air quality and temperature were sampled at 22 monitoring stations between 12:00 a.m. January 1, 2008 and 11:00 p.m. December 31, 2012, for more than 38,000 observations.^{30,31} Ozone was sampled at 19 of the stations and particulate matter was sampled at 6 of the stations.

Statistical Analysis

Hotspot Analysis

Spatial clustering of emergency room visits for acute bronchitis was examined using the local Getis-Ord G_i^* statistic, similarly to Beck et al.¹ The G_i^* statistic was used to identify ZIP codes with higher than expected rates of emergency room visits, assuming a random spatial distribution. Analysis was performed using Esri ArcGIS.^{32,33}

Analysis of Socioeconomic Factors and Emergency Room Visit Rate Variation

ZIP codes were categorized by quintile of emergency room visit rate, and differences in socioeconomic status (SES) variables were analyzed following the procedure described by Beck et al. Differences between the distribution of SES variable medians across emergency room visit rate quintiles were assessed using the nonparametric Kruskal-Wallis test.³⁴ Variables were selected from the 2010 American Community Survey (ACS 2010) to allow comparison to previous studies and to highlight vulnerable populations. Variables selected included: adults with less than a high school education, unemployment, median annual household income, vacant homes, renter-occupied homes, households without a vehicle, median home value, median rent, residents never married, disabled residents and households with gross rent greater than 35% of income. We calculated the percent of older adults and children with no health insurance as the proportion of children ages 0–19 and adults 65 and older with no insurance using ACS 2010 data. To allow comparison of our findings to Beck et al., we performed a sensitivity analysis that found that the results of our analysis of SES effects do not differ when the analysis is limited to emergency room visits by patients age 21 years or younger.

Independence of Air Quality and SES Predictors

To assess the independence of air quality and SES for confounding relationships, average air quality for the study period was compared between socioeconomic quintiles using a Kruskal-Wallis test. We also calculated the correlation between daily air quality predictors and SES quintile as well as the correlation among air quality predictors.

Mixed Effects Modeling

We created four Generalized Linear Mixed Models (GLMMs) with Poisson log-link functions to examine the association between air quality, socioeconomic status and emergency room visits for acute bronchitis. The response variable was the daily rate of emergency room visits per 100,000 residents in each of 32 ZIP codes. ZIP codes with no emergency room visits during the study period were excluded. Similarly to Berhane et al. our models incorporate a hierarchical structure of predictors at the day and ZIP code level.³⁵ Fixed effects for air quality predictors were included in all GLMMs. The air quality predictors included lagged daily mean PM10, PM2.5, Ozone and Temperature for each ZIP code. ZIP code level predictors included random spatial effects and socioeconomic effects.

Fixed Effects

Air quality and temperature were interpolated at the geographic center of each ZIP code using covariance kriging.^{36–38} Empirical variograms were calculated using hourly observations. Seven potential models were fit to select

the most appropriate kriging functions. A Gaussian model was selected for zone and PM_{2.5}, a Cauchy model was selected for PM₁₀ and an exponential model was selected for temperature. The kriging functions were applied to interpolate hourly estimates of air quality for each ZIP code. The hourly estimates were aggregated into time series of 24-hour mean measurements.³⁹

Methods described by Schwartz and others were used for seasonal trend adjustment and smoothing.^{40–42} In order to correct for dominance of seasonal trends in the GLMM and focus the analysis on changes in daily air quality, harmonic models were fit to de-trend air quality and temperature predictors and emergency room visits in each ZIP code.^{40,43} Gaussian Kernel smoothing (window width of ten days) was applied to the emergency room visit series. To determine the appropriate lag periods for predictors, we adapted methods described by Katsouyanni et al. to account for variation between ZIP codes.⁴² Cross Correlation Functions (CCFs) were fit to determine the lag period with strongest correlation between each predictor variable and emergency room visits for acute bronchitis in each ZIP code.³⁹ A histogram of strongest correlation lags from all ZIP codes, weighted by emergency room visits in each ZIP code, was used to choose a single lag period for each variable to be used in regression modeling. We performed a sensitivity analysis that found little difference in the conclusions of models fitted with the same lag period for all ZIP codes and a lag period individualized for each ZIP code. Lagged time series were created for Ozone (k = 3 days), PM₁₀ (k = 10 days), PM_{2.5} (k = 4 days) and Temperature (k = 5 days).⁴⁴

Random Effects

A base GLMM was created to model the effects of air quality on emergency room visits for acute bronchitis. Random effects were included to account for potential spatial correlation and observation imbalance between ZIP codes.⁴⁵ Random intercepts and slopes for lagged PM₁₀, PM_{2.5}, Ozone and Temperature were fit by ZIP code.

Three additional GLMMs were fit by adding effects for ZIP code socioeconomic status to the base model. Factor analysis was used to classify ZIP codes in socioeconomic quintiles.⁴⁶ A single socioeconomic factor was created by incorporating the most significant variables from the emergency room visit quintile analysis, which included: vacant homes, median home value, median rent and adults and children with no health insurance. The second and third GLMMs were fit by subsequently adding random intercepts by socioeconomic quintile and random slopes for air quality predictors by socioeconomic quintile. A fourth GLMM was fit by adding fixed effect for socioeconomic quintile to the base model of air quality predictors.

Computation of GLMMs

Markov-Chain Monte Carlo (MCMC) estimation was used to fit the mixed effects Poisson GLMs because of possible zero inflation or spatial autocorrelation in the emergency room visit rate distribution.⁴⁵ Three MCMC chains were computed for each model, using a parameter expanded proper Cauchy prior.^{47,48} The base air quality model was run for 80,000 iterations with a burn-in of 30,000 and a step of 50. The socioeconomic status models were run for 120,000 iterations with the same burn-in and step. Convergence was tested using the Gelman-Rubin statistic.^{49,50} Gelman-Rubin statistic values less than 1.02 are considered indicative of good convergence as a test for model reliability. Models were assessed by comparing the variances of the random effects and quantifying fit with the Deviance Information Criterion, a measure similar to the AIC for Bayesian mixed models.^{45,51}

Results

Emergency Room Visits for Acute Bronchitis

5,620 emergency room visits with a primary ICD-9 diagnosis of 4660 by residents of San Luis Obispo and Santa Barbara counties were reported during the study period. 1,352 cases were residents over age 55; 1,667 cases were residents age 22–54; 2,601 cases were residents age 0–21 and 3,305 cases were women. Women 21 and older visited the emergency room more frequently than men (2,487 men vs. 1,466 women). The total daily emergency room visits for acute bronchitis for the study area had a mean of 3.84 and followed a Poisson distribution, with 0–16 total visits per day and 0–2 visits per day in any single ZIP code. At least one emergency room visit was reported on more than 96% of days in the study period. The rate of acute bronchitis emergency room visits showed strong seasonal periodicity. Visits peaked in the late winter between December and February. The lowest rates of admission were seen in midsummer.

Hotspot Analysis

Geographic variation was observed in emergency room visit rates. The emergency room visits included residents of 32 of the 49 ZIP codes in the study area. The four-year rate of emergency room visits for acute bronchitis is between 2 and 17 visits per 1,000 residents for all ZIP codes. Two clusters of ZIP codes with high emergency room visit rates were observed around the communities of Templeton and Lompoc, CA (**Figure 1**). ZIP codes 93465, 93432 and 93454 in these clusters were found to be hotspots by the G_i^* analysis at 90% significance (**Figure 2**). No significant cold spots were found, but lower emergency room visit rates were observed in San Luis Obispo County along the coast.

Socioeconomic Differences across Emergency Room Visit Rate Quintiles

The median rate for acute bronchitis emergency room visits varied significantly across quintiles. 0.0, 0.0, 4.28, 8.07 and 13.5 emergency room visits were reported per 1,000 residents over four years in the lowest to highest quintiles, respectively ($p < .001$). The lowest and low medium quintiles were combined during analysis because of indistinguishable rates of 0.0 emergency room visits during the study period. Significant differences in socioeconomic measures were observed among quintiles (**Table 2**). Median home value and median rent were significant ($p = .003$ and $p < .001$, respectively) with similar trends of lower home values and rent in higher quintiles of emergency room visit rates, with the exception of the lowest median home values and rent being in the combined low and low medium emergency room visit quintiles. The percentage of vacant homes and the percentage of older adults and children without health insurance were close to significant ($p = .036$ and $p = .083$, respectively).

Generalized Linear Mixed Models

The MCMC estimates for the mixed effects models showed strong convergence with a Gelman-Rubin statistic of less than 1.01 and low autocorrelation. Posterior predictive tests for model misspecification found only slight zero-inflation.

No significant difference was found in the distribution of air quality between socioeconomic quintiles (KW p -values ≥ 0.20). No meaningful correlations between predictor variables were found. Air quality predictors were found to be correlated with SES quintile with Pearson's R coefficients less than 0.0001. Air quality predictors were

found to be correlated with other air quality predictors with a Pearson's R coefficient less than 0.32.

Lagged daily mean PM10 concentration was found to be a significant predictor of emergency room visits in models including only air quality predictors and random effects for spatial autocorrelation (see model A in **Table 3**).

The addition of a fixed affect for SES quintile was not found to improve the fit of the model (see model D in **Table 3**). The DIC did not decrease meaningfully (5%) compared to the base model, and the credible interval of the SES quintile effect was wide and correlation was observed in most posterior distributions, indicating poor fit.

The addition of random intercepts for SES quintile improved the fit of the model leading to a large reduction of the DIC (41%) compared to the base model (see model B in **Table 3**). SES quintiles 1–3 generally have larger intercepts than SES quintiles 4–5, suggesting that lower SES may be associated with higher rates of emergency room visits. Lagged daily average temperature was a significant predictor in the random intercepts model.

The addition of random slopes for SES quintile resulted in a minimal decrease of the DIC (<1%) compared to the random intercepts model (see model C in **Table 3**). This model was considered to be a better fit because there were only minor changes in ZIP code random effect variances from the random intercepts model, and the most of the variance in the SES quintile random intercepts was partitioned into the SES quintile random slopes. No air quality fixed effects were found to be significant predictors of emergency room visits. While the variances of the SES quintile random slopes are objectively close to zero, the interquartile range size of the effects is 2% to 96% of the size of the average slopes. There is no consistent association between SES quintile and the size of the random slope for individual quintiles, but the mean effect of quintiles 1 and 2 is larger than the mean effect of quintiles 4 and 5 for all variables. However, the 95% credible intervals from the posterior distributions for all SES quintile random slopes overlap zero. This suggests that there may be an association between low SES and increased effect of air quality that this sample does not have enough power to identify.

Discussion

Findings

Our Hotspot analysis provides clear evidence that significant variation exists in emergency room visit rates for acute bronchitis across ZIP codes in rural counties with small to medium size cities. There is also clear evidence that variation in emergency room visit rates is associated with significant differences in ZIP code level socioeconomic factors, including median home value and median rent.

Relationship to Previous Literature

However, we did not find significant differences for all socioeconomic factors found to be associated with emergency room visit rates by Beck et al. in Cincinnati, OH. We are unable to conclude that SES modifies the short-term effect of air quality on emergency room visit rates for acute bronchitis. Although we observe differences in the random coefficients of air quality predictors on emergency room visit rates that suggest lower SES is associated with greater effects of air quality, these differences do not appear large enough to be statistically meaningful.

Limitations

Since our emergency room visit dataset time series are very sparse, we believe this study is underpowered for the effect sizes of air quality and SES. We anticipate that a longer study period or a similar investigation of a more

prevalent respiratory illness, such as asthma, would result in more conclusive results. Previous research has found significant short term effects of air quality during summer months only.^{2,9} Our analysis included data from all seasons, and it is not known how seasonal limitations (such as limiting the analysis to flu season, when acute bronchitis is most prevalent) could have affected the conclusions. Like all ecological studies, our analysis is limited by the lack of individual level data about emergency room visits, including patient identity, socioeconomic status and confounding health factors.⁵² As noted previously, emergency room visits do not capture the full burden of acute bronchitis, especially cases that result in private medical care.¹ Observational studies of emergency room visits may also be confounded by varying rates of emergency room utilization, which have been shown to be higher among individuals with lower SES.

Significance

These results show significant variation in the epidemiology of respiratory infection rates at a small geographic scale in rural areas and small cities, with implications for health care, policy and research. While counties with low overall rates of respiratory infections may not appear to require control efforts, hotspots of disease can exist, potentially in places where vulnerable low socioeconomic status populations live. Therefore, public health agencies should consider small-scale geospatial analysis as one method to identify disparities in health and efficiently deliver resources to the populations most at risk.

We found strong evidence that socioeconomic factors at a small geographic scale should be considered important in epidemiological studies of respiratory infections in small cities and rural communities. While socioeconomic factors are associated with differences in rate of emergency room visits for respiratory infections in both large urban cities and rural areas, a different set of socioeconomic factors may be important in rural areas, and the relationships with disease may be different. Similarly, the effect of air quality on emergency room visits for respiratory infections may vary considerably between nearby communities and may be modified by socioeconomic factors. Analyses that ignore these relationships risk biased conclusions. Our suggestions for research and health policy for small cities and rural areas reinforce existing recommendations developed using data from large metropolitan areas.⁵³

Conclusion

We found clear evidence of significant variation in emergency room visit rates for acute bronchitis at a small geographic scale in rural counties with small to medium size cities. Variation in emergency room visit rates was associated with significant differences in socioeconomic factors including median home value and median rent.

References

1. Beck, A.F., Florin, T.A., Campanella, S., & Shah, S.S. Geographic variation in hospitalization for lower respiratory tract infections across one county. (2015). *JAMA Pediatr*, **169**, 846–854.
2. Tager, Ira et al. Fresno Asthmatic Children's Environment Study (FACES) Final Report. (California Air Review Board, 2006).
3. Mutlu, G. M., Bellmeyer, A., & Budinger, G. R. S. (2006). Air pollution impairs lung's ability to clear edema fluid. *Am. J. Cardiol.* **98**, 423–424.
4. Burnett, R. T. et al. (2001). Association between ozone and hospitalization for acute respiratory diseases in children less than 2 years of age. *Am. J. Epidemiol.* **153**, 444–452.
5. Pino, P., Walter, T., Oyarzun, M., Villegas, R., & Romieu, I. (2004). Fine particulate matter and wheezing illnesses in the first year of life. *Epidemiology* **15**, 702–708.
6. Bell ML, M. A. (2004). Ozone and short-term mortality in 95 US urban communities, 1987-2000. *JAMA* **292**, 2372–2378.
7. Davidson, C. I., Phalen, R. F., & Solomon, P. A. (2005). Airborne particulate matter and human health: A review. *Aerosol Sci. Technol.* **39**, 737–749.
8. Samet, J. M., Dominici, F., Curriero, F. C., Coursac, I., & Zeger, S. L. (2000). Fine particulate air pollution and mortality in 20 U.S. cities, 1987–1994. *N. Engl. J. Med.* **343**, 1742–1749.
9. Ostro, B. (1995). Fine particulate air pollution and mortality in two Southern California counties. *Environ. Res.* **70**, 98–104.
10. Anderson, J. O., Thundiyil, J. G., & Stolbach, A. (2012). Clearing the air: a review of the effects of particulate matter air pollution on human health. *J. Med. Toxicol. Off. J. Am. Coll. Med. Toxicol.* **8**, 166–175.
11. Eftim, S. E., Samet, J. M., Janes, H., McDermott, A., & Dominici, F. (2008). Fine particulate matter and mortality. *Epidemiology* **19**, 209–216.
12. Romeo, E. et al. (2006). PM 10 exposure and asthma exacerbations in pediatric age: a meta-analysis of panel and time-series studies. *Epidemiol. Prev.* **30**, 245–254.
13. Anderson, H. R., Bremner, S. A., Atkinson, R. W., Harrison, R. M., & Walters, S. (2001). Particulate matter and daily mortality and hospital admissions in the west midlands conurbation of the United Kingdom: associations with fine and coarse particles, black smoke and sulphate. *Occup. Environ. Med.* **58**, 504–510.
14. Atkinson, R. W. et al. (2001). Acute effects of particulate air pollution on respiratory admissions results from APHEA 2 project. *Am. J. Respir. Crit. Care Med.* **164**, 1860–1866.

15. Zanobetti, A. et al. (2003). The temporal pattern of respiratory and heart disease mortality in response to air pollution. *Environ. Health Perspect.* **111**, 1188.
16. Darrow, L. A. et al. (2014). Air Pollution and Acute Respiratory Infections Among Children 0–4 Years of Age: An 18-Year Time-Series Study. *Am. J. Epidemiol.* **180**, 968–977.
17. Simoni, M. et al. (2015). Adverse effects of outdoor pollution in the elderly. *J. Thorac. Dis.* **7**, 34–45.
18. Bentayeb, M. et al. (2012). Adverse respiratory effects of outdoor air pollution in the elderly. *Int. J. Tuberc. Lung Dis. Off. J. Int. Union Tuberc. Lung Dis.* **16**, 1149–1161.
19. Sung, R. Y., Chan, R. C., Tam, J. S., Cheng, A. F., & Murray, H. G. (1992). Epidemiology and aetiology of acute bronchiolitis in Hong Kong infants. *Epidemiol. Infect.* **108**, 147–154.
20. Guo, L. J., Zhao, A., Chen, R. J., Kan, H. D., & Kuang, X. Y. (2014). Association between ambient air pollution and outpatient visits for acute bronchitis in a Chinese city. *Biomed. Environ. Sci.* **27**, 833–840.
21. Ostro, B., Roth, L., Malig, B., & Marty, M. (2009). The effects of fine particle components on respiratory hospital admissions in children. *Environ. Health Perspect.* **117**, 475–480.
22. Martins, M. C. H. et al. (2003). Influence of socioeconomic conditions on air pollution adverse health effects in elderly people: an analysis of six regions in São Paulo, Brazil. *J. Epidemiol. Community Health* **58**, 41.
23. Forastiere, F. et al. (2007). Socioeconomic status, particulate air pollution, and daily mortality: Differential exposure or differential susceptibility. *Am. J. Ind. Med.* **50**, 208–216.
24. Jerrett, M. et al. (2003). Do socioeconomic characteristics modify the short term association between air pollution and mortality? Evidence from a zonal time series in Hamilton, Canada. *J. Epidemiol. Community Health* **58**, 31.
25. International Classification of Diseases: Ninth Revision (ICD-9) | *Annals of Internal Medicine* | American College of Physicians. Available at: <http://annals.org/aim/article/691852/international-classification-diseases-ninth-revision-icd-9>. (Accessed: 16th November 2016)
26. Office of Statewide Health Planning and Development. Emergency Department & Ambulatory Surgery Data. Ca.gov (2016). Available at: <http://www.oshpd.ca.gov/HID/Products/EmerDeptData/>. (Accessed: 22nd March 2016)
27. United States Census Bureau. American FactFinder. Available at: <http://factfinder.census.gov/faces/nav/jsf/pages/searchresults.xhtml?refresh=t>. (Accessed: 22nd April 2016)
28. US EPA, O. AirData website File Download page. EPA.gov Available at: http://aqsdr1.epa.gov/aqsweb/aqstmp/airdata/download_files.html#Raw. (Accessed: 22nd March 2016)
29. CA Air Resources Board. Meteorology Data Query Tool (PST). arb.ca.gov (2011). Available at: <http://www.arb.ca.gov/aqmis2/metslect.php>. (Accessed: 22nd March 2016)

30. Annual Air Quality Report. Santa Barbara County Air Pollution Control District Available at: <http://www.ourair.org/sbc/annual-air-quality-report/>. (Accessed: 22nd March 2016)
31. Air Quality Reports. San Luis Obispo County Air Pollution Control District Available at: <http://www.slocleanair.org/air/annualreport.php>. (Accessed: 22nd March 2016)
32. Esri - GIS Mapping Software, Solutions, Services, Map Apps, and Data. Available at: <http://www.esri.com/>. (Accessed: 22nd March 2016)
33. ESRI, inc. ArcGIS - USA ZIP Code Areas. (2016). Available at: <http://www.arcgis.com/home/item.html?id=8d2012a2016e484dafaac0451f9aea24>. (Accessed: 22nd March 2016)
34. R Core Team. R: A language and environment for statistical computing. (R Foundation for Statistical Computing, 2015).
35. Berhane, K., Gauderman, W. J., Stram, D. O., & Thomas, D. C. (2004). Statistical issues in studies of the long-term effects of air pollution: The Southern California children's health study. *Stat. Sci.* **19**, 414–434.
36. Wong, D. W., Yuan, L., & Perlin, S. A. (2004). Comparison of spatial interpolation methods for the estimation of air quality data. *J. Expo. Sci. Environ. Epidemiol.* **14**, 404–415.
37. Ribeiro Jr, P. J. & Diggle, P. J. (2001). geoR: a package for geostatistical analysis. *R News* **1**, 14–18.
38. Liu, L.-J. S. & Rossini, A. J. (1996). Use of kriging models to predict 12-hour mean ozone concentrations in metropolitan Toronto—a pilot study. *Environ. Int.* **22**, 677–692.
39. Zeileis, A. & Grothendieck, G. (2005). zoo: S3 infrastructure for regular and irregular time series. *ArXiv Prepr. Math0505527*.
40. Bhaskaran, K., Gasparrini, A., Hajat, S., Smeeth, L., & Armstrong, B. (2013). Time series regression studies in environmental epidemiology. *Int. J. Epidemiol.* **42**, 1187–1195.
41. Schwartz, J. et al. (1996). Methodological issues in studies of air pollution and daily counts of deaths or hospital admissions. *J. Epidemiol. Community Health* **50**, S3–11.
42. Katsouyanni, K. et al. (1996). Short term effects of air pollution on health: A European approach using epidemiologic time series data: the APHEA protocol. *J. Epidemiol. Community Health* **50**, S12–S18.
43. Lück, S., Thurley, K., Thaben, P. F., & Westermark, P. O. (2016). Rhythmic degradation explains and unifies circadian transcriptome and proteome data. *Cell Rep.* **9**, 741–751 (2014).
44. Zeileis, A. dynlm: Dynamic Linear Regression. (2016).
45. Hadfield, J. D. & others. (2010). MCMC methods for multi-response generalized linear mixed models: the MCMCglmm R package. *J. Stat. Softw.* **33**, 1–22.

46. Revelle, W. (2016). *psych: Procedures for Psychological, Psychometric, and Personality Research*.
47. Gelman, A. & others. (2006). Prior distributions for variance parameters in hierarchical models (comment on article by Browne and Draper). *Bayesian Anal.* **1**, 515–534.
48. Barr, D. J., Levy, R., Scheepers, C., & Tily, H. J. (2013). Random effects structure for confirmatory hypothesis testing: Keep it maximal. *J. Mem. Lang.* **68**.
49. Gelman, A., & Rubin, D. B. (1992). Inference from iterative simulation using multiple sequences. *Stat. Sci.* **7**, 457–472.
50. Plummer, M., Best, N., Cowles, K., & Vines, K. (2006). CODA: Convergence diagnosis and output analysis for MCMC. *R News* **6**, 7–11.
51. Zhu, L. & Carlin, B. P. (2000). Comparing hierarchical models for spatio-temporally misaligned data using the deviance information criterion. *Stat. Med.* **19**, 2265–2278.
52. Kunzli, N., & Tager, I. B. (1997). The semi-individual study in air pollution epidemiology: A valid design as compared to ecologic studies. *Environ. Health Perspect.* **105**, 1078–1083.
53. Hubbell, B., Fann, N., & Levy, J. I. (2009). Methodological considerations in developing local-scale health impact assessments: balancing national, regional, and local data. *Air Qual. Atmosphere Health* **2**, 99–110.

Tables and Figures

| Name | Total | ZIP Codes in Study Area (n=49) | | |
|------------------------------------|--------------------|--------------------------------|--------------------|----------------------|
| | | Mean | Median | SD |
| Population | 765836 | 15316.72 | 9971 | 15635.31 |
| Male | 392662 (51.6%) | 7853.24 (51.6%) | 4806 (50.41%) | 8055.47 (5.68%) |
| White | 564137 (76.13%) | 11282.74 (76.13%) | 7593.5 (78.63%) | 10961.04 (15.16%) |
| Black/ African American | 16638 (1.74%) | 332.76 (1.74%) | 78.5 (1.02%) | 609.94 (2.21%) |
| Asian | 30077 (2.77%) | 601.54 (2.77%) | 205.5 (2.34%) | 1013.22 (2.11%) |
| Two or more races | 32448 (3.85%) | 648.96 (3.85%) | 317 (3.66%) | 777.48 (1.49%) |
| Other | 113410 (14.31%) | 2268.2 (14.31%) | 664 (9.64%) | 3558.22 (14.81%) |
| Hispanic or Latino | 271721 (32.28%) | 5434.42 (32.28%) | 1977.5 (27.82%) | 8149.06 (21.13%) |
| 0-19 | 200532 (25.18%) | 4010.64 (25.18%) | 573.5 (21.11%) | 4725.76 (6.95%) |
| 65+ | 108997 (15.75%) | 2179.94 (15.75%) | 276.25 (12.18%) | 2042.35 (6.64%) |
| Population Density (n/ sqmi) | 748.99 | 748.99 | 22.49 | 1665.88 |

Table 1: Sample Characteristics

| Name | Q1 & Q2 | Q3 | Q4 | Q5 | P |
|---|-----------|-----------|-----------|-----------|-----------|
| ER visits per 1,000 (n/year) | 0.0 | 4.28 | 8.07 | 13.85 | <0.001*** |
| Adults with less than a high school education (%) | 0.85 | 0.90 | 0.87 | 0.87 | 0.192 |
| Unemployment Rate (%) | 0.05 | 0.03 | 0.04 | 0.04 | 0.219 |
| Median Annual Household Income (\$) | 55574.50 | 71357.50 | 61074.50 | 58670.00 | 0.123 |
| Vacant Homes (%) | 17.90 | 10.20 | 6.20 | 9.00 | 0.036* |
| Renter Occupied Homes (%) | 37.85 | 41.20 | 47.45 | 37.75 | 0.834 |
| Households that do not own a vehicle (%) | 3.40 | 4.60 | 5.80 | 4.65 | 0.367 |
| Median Home Value (\$) | 243500.00 | 665000.00 | 662750.00 | 371500.00 | 0.003** |
| Median Rent (\$) | 1033.50 | 1407.00 | 1364.50 | 1102.50 | <0.001*** |
| Residents Never Married (%) | 34.65 | 38.35 | 33.90 | 29.20 | 0.338 |
| Disabled Residents (%) | 0.97 | 0.99 | 1.00 | 0.97 | 0.798 |
| Gross Rent greater than 35% household income (%) | 48.00 | 49.30 | 46.35 | 46.75 | 0.731 |
| Percentage of older (65+) adults and children (0-19) with no health insurance (%) | 1.18 | 0.97 | 0.97 | 1.27 | 0.083* |

Legend: * $p < .10$ (close to significant), ** $p < .01$ (significant), *** $p < .001$ (significant)

Table 2: Socioeconomic Factors in Emergency Room Visit Rate Quintiles

| Model | A. Base Model | | | B. Random Intercepts by SES Quintile | | | C. Random Slopes by SES Quintile | | | D. SES Quintile Fixed Effect | | |
|-------------------------|--------------------|------------------------|-----------|--------------------------------------|------------------------|-----------|----------------------------------|------------------------|-----------|------------------------------|------------------------|-----------|
| DIC | 679769.5 | | | 403512.9 | | | 403457.7 | | | 644410.7 | | |
| Fixed Effects | Posterior Mode | Coefficient Percentile | | Posterior Mode | Coefficient Percentile | | Posterior Mode | Coefficient Percentile | | Posterior Mode | Coefficient Percentile | |
| | | 2.50 % | 97.50 % | | 2.50 % | 97.50 % | | 2.50 % | 97.50 % | | 2.50 % | 97.50 % |
| Intercept | -0.1055 | -0.7181 | 0.6042 | -2.0545 | -3.4346 | -0.8493 | -2.2485 | -3.4929 | -1.0117 | 0.8471 | -1.3503 | 2.7012 |
| Lag_Temp | -0.0487 | -0.1251 | 0.0695 | -0.0654 | -0.0934 | -0.0359 | -0.0667 | -0.1740 | 0.0556 | -0.0315 | -0.1987 | 0.1715 |
| Lag_PMten | 0.0036 | 0.0006 | 0.0076 | 0.0110 | -0.0149 | 0.0293 | -0.0092 | -0.1065 | 0.1328 | 0.0040 | -0.0002 | 0.0084 |
| Lag_PMTwo | -0.0068 | -0.0223 | 0.0093 | -0.0013 | -0.0420 | 0.0250 | -0.0136 | -0.1333 | 0.1042 | -0.0044 | -0.0261 | 0.0158 |
| Lag_Ozone | 0.7611 | -12.3018 | 9.1841 | -5.6422 | -23.0301 | 14.9378 | -1.8456 | -21.3271 | 14.5221 | -2.9438 | -12.6263 | 9.9823 |
| SES quintile | | | | | | | | | | -0.1937 | -0.7679 | 0.4520 |
| Random Effects | Posterior Variance | Coefficient Percentile | | Posterior Variance | Coefficient Percentile | | Posterior Variance | Coefficient Percentile | | Posterior Variance | Coefficient Percentile | |
| | | 2.50 % | 97.50 % | | 97.50 % | 2.50 % | | 2.50 % | 97.50 % | | 2.50 % | 97.50 % |
| units | 1.0115 | 1.0115 | 1.0115 | 12.5055 | 12.3306 | 12.6762 | 12.4984 | 12.3157 | 12.6795 | 0.9028 | 0.9028 | 0.9028 |
| <i>By ZIP Code:</i> | | | | | | | | | | | | |
| Intercept | 3.4584 | 2.0362 | 5.1463 | 11.3518 | 6.4266 | 17.4310 | 11.2774 | 6.3908 | 17.6967 | 3.6377 | 1.8666 | 5.6191 |
| Lag_Temp | 0.0510 | 0.0056 | 0.1445 | 0.0062 | 0.0035 | 0.0094 | 0.0066 | 0.0036 | 0.0102 | 0.1207 | 0.0096 | 0.2768 |
| Lag_PMten | 0.0001 | 0.0000 | 0.0001 | 0.0039 | 0.0023 | 0.0058 | 0.0043 | 0.0024 | 0.0067 | 0.0006 | 0.0000 | 0.0002 |
| Lag_PMTwo | 0.0022 | 0.0011 | 0.0033 | 0.0082 | 0.0047 | 0.0125 | 0.0091 | 0.0046 | 0.0137 | 0.0040 | 0.0010 | 0.0051 |
| Lag_Ozone | 913.0605 | 549.6501 | 1392.8088 | 2471.8034 | 1343.4210 | 3846.1064 | 2481.2271 | 1331.7519 | 3785.6584 | 933.0524 | 525.1977 | 1412.2857 |
| <i>By SES Quintile:</i> | | | | | | | | | | | | |
| Intercept | | | | 0.0980 | 0.0024 | 0.3586 | 0.0293 | 0.0057 | 0.0730 | | | |
| Lag_Temp | | | | | | | 0.0156 | 0.0052 | 0.0315 | | | |
| Lag_PMten | | | | | | | 0.0156 | 0.0046 | 0.0323 | | | |
| Lag_PMTwo | | | | | | | 0.0160 | 0.0042 | 0.0322 | | | |
| Lag_Ozone | | | | | | | 0.0306 | 0.0047 | 0.0829 | | | |

Table 3: Generalized Linear Mixed Effects Models

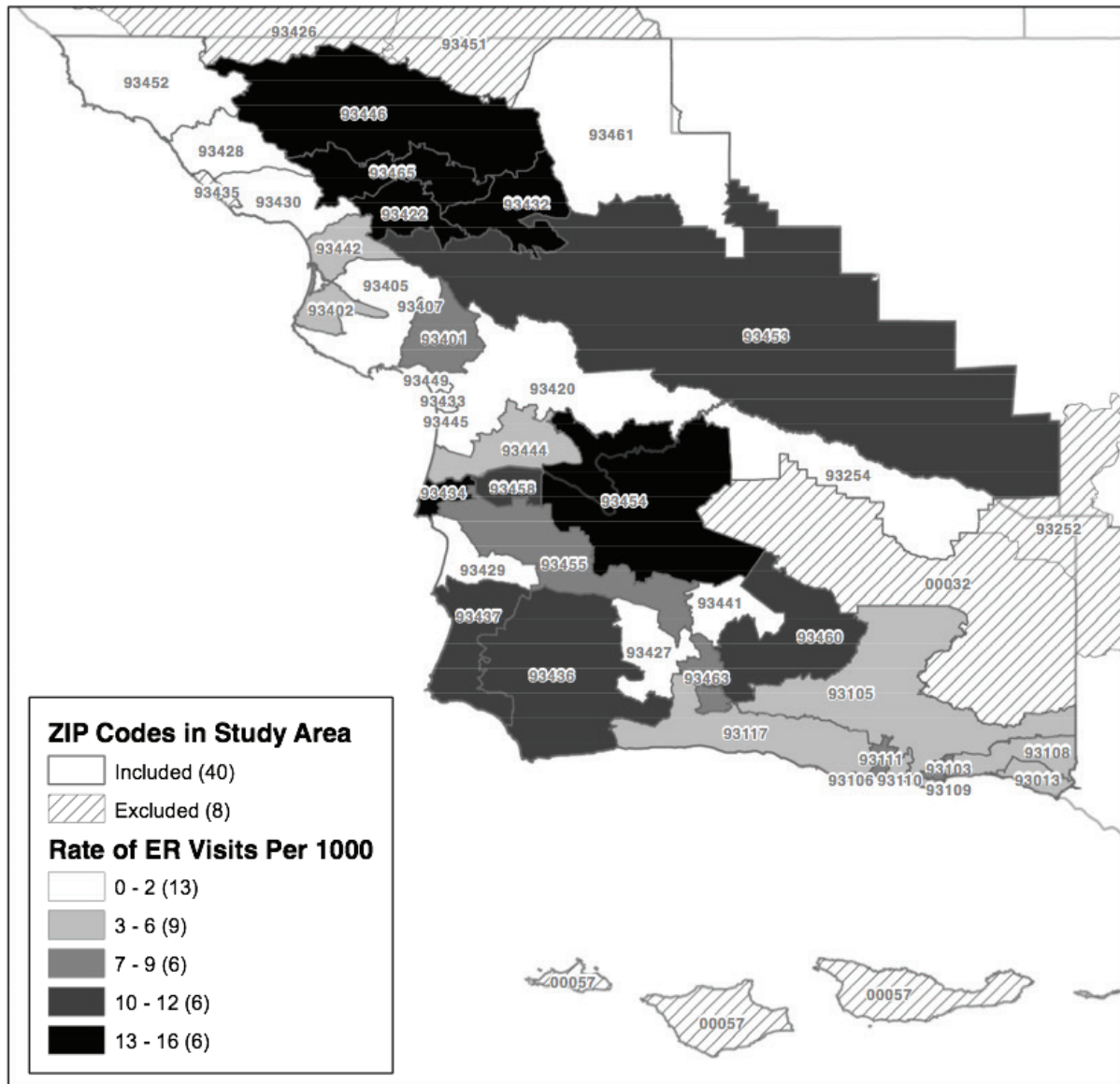


Figure 1: Emergency Room Visits for Acute Bronchitis

Map of San Luis Obispo and Santa Barbara counties with a choropleth of 4-year rate (2009–2012) of ER visits for acute bronchitis. ZIP codes not fully contained in San Luis Obispo or Santa Barbara county or with a population less than 75 were excluded from the study area.

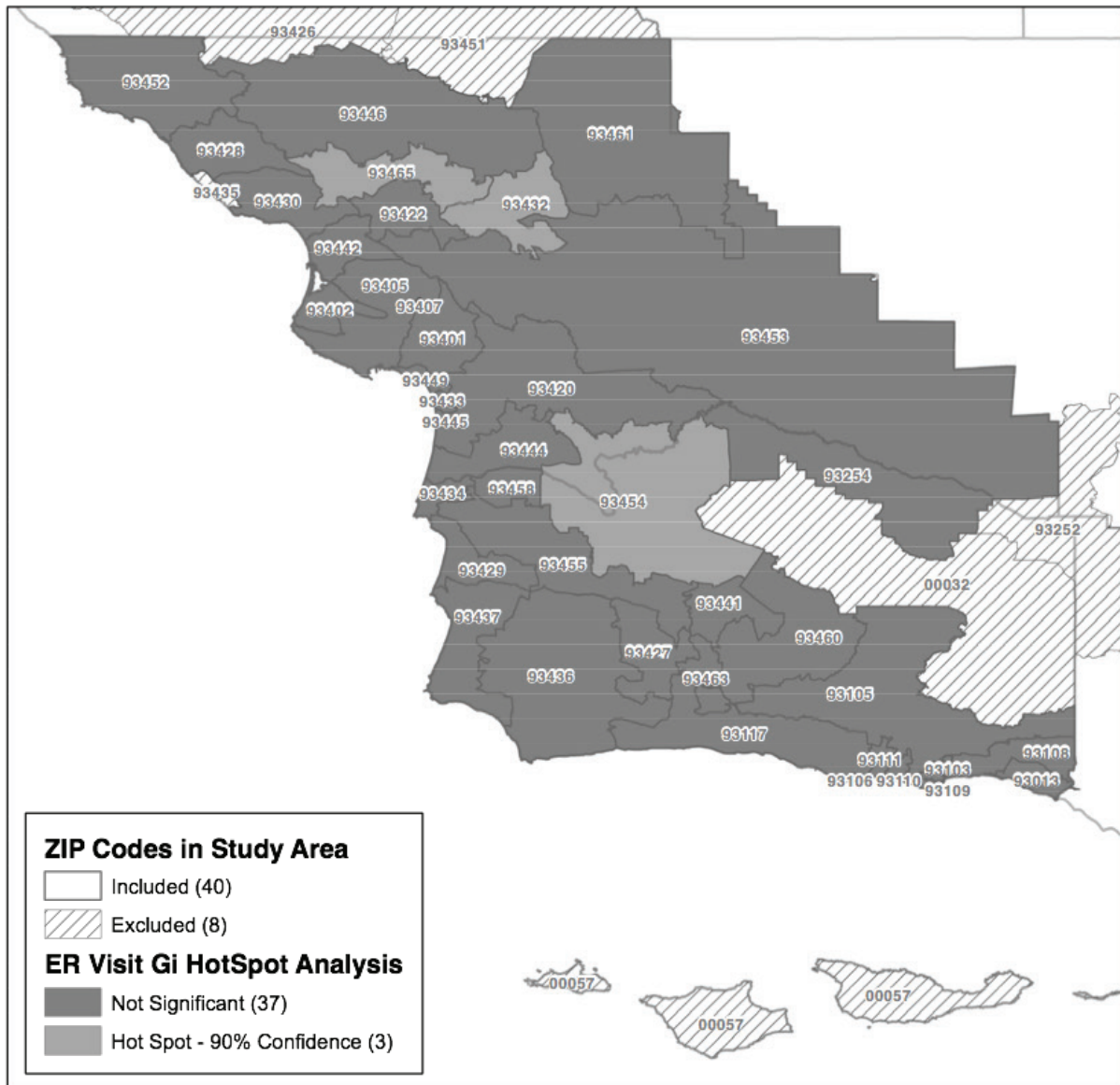


Figure 2: Hotspot Analysis

Map of San Luis Obispo and Santa Barbara counties with Hotspots of ER visits for acute bronchitis highlighted. Hotspots were identified using the Getis-Ord G_i^* statistic calculated using the four-year rate of ER visits with Esri ArcGIS software.

About the Author

Sean Lang-Brown graduated from California Polytechnic State University, San Luis Obispo in 2016 with a B.S. in Biology and a minor in Statistics. Sean was a participant in the Cal Poly Honors Program, an intern at the San Luis Obispo County Public Health Department Tobacco Control Program, and a founding member of the Cal Poly Public Health Club. He was also treasurer of Yo Tango (the Cal Poly Argentine Tango Club) and involved in the Cal Poly Ballroom Dance Club. Sean is now a Clinical Research Coordinator in the Division of Geriatrics at The University of California, San Francisco.

Acknowledgements

This work was funded by a grant from the US Green Building Council and Summer Research Funding from the Cal Poly College of Science and Mathematics. The researchers utilized campus resources including the Kennedy Library GIS lab and DIRAC high performance computing cluster. Sean Lang-Brown was responsible for all aspects of this study. Heather Starnes and Gary Hughes provided advice for study design and analysis. The researchers would like to thank Ann McDowell (SLO County Public Health Department), Stacy White (US Green Building Council), Aeron Arlin-Genet (SLO Air Pollution Control District), Gary Arcemont (SLO Air Pollution Control District), Kelly Main (Cal Poly) and Jennifer Franich (Cal Poly) for their support.

II.

Interdisciplinary Research



Radiation Exposure During Space Travel: Using Radioisotopes for a Comparative Study of Human Feces and Urine as Integrated Shield Components

By Noah Falck

Abstract

Earth's atmosphere and magnetosphere deflect and absorb the majority of harmful radiation traversing space; however, once outside Low Earth Orbit (LEO), payloads are exposed to Galactic Cosmic Rays (GCR) and Solar Particle Events (SPE). While humans possess capabilities that hardware does not, we are uniquely vulnerable to radiation. Detrimental effects range from nausea and dizziness brought by intense, short-term doses to increased cancer risk and impaired cognitive function associated with chronic exposure. This paper aims to explore the use of human waste (feces and urine) as radiation shields in a comparative study of urine vs. water and hydrated vs. dehydrated feces. GCRs contain particles with energy orders of higher magnitude than SPEs, which makes them impractical to shield against. Conversely, SPEs occur with higher frequency and at a lower level that is practically attenuated. To this end, an SPE surrogate was validated and exposure (Counts) on the leeward side of the respective shields was measured. Counts per Minute (CPM) were obtained by applying a multiplicative factor to Counts. As expected, CPM behind a urine shield did not differ from CPM behind a water shield (t-test, $p < 0.05$). Similarly, no difference in leeward CPM was found between hydrated and dehydrated feces shields (t-test, $p < 0.05$). The lack of differences between water and urine may be a result of urine being primarily composed of water. While fecal matter is made primarily of water as well, the solid content likely masks the loss of water's attenuation properties.

Acronyms

| | |
|-----|----------------------------|
| CPM | Counts per Minute |
| GCR | Galactic Cosmic Ray |
| GMW | Geiger-Mueller tube window |

| | |
|---------|--|
| HIMAC | Heavy Ion Medical Accelerator in Chiba |
| LEO | Low Earth Orbit |
| OLTARIS | Online Tool for Assessment of Radiation in Space |
| RI | Radioisotope |
| SPE | Solar Particle Event |

Introduction

As humans continue to travel farther from Earth to explore space, we encounter a Pandora's box of biological hazards. Chief among dangers to *Homo sapiens* is the damage caused by radiation (Horneck et al., 2006), which can be classified into two main categories: acute and chronic symptoms. Acute symptoms include nausea and vomiting and are typically the result of exposure to a high dose during a short timespan. Chronic symptoms vary from increased cancer risk to cataract development (Langford, n.d.) and are associated with prolonged exposure to radiation. With space travel, there are two 1,000 lb gorillas in the room: mass and money. This paper will involve the former. The mass limitation requires that waste is minimized and that many materials serve multiple purposes, such as the use of drinking water as part of a Solar Particle Event (SPE) shelter (Simon, Clowdsley, & Walker, n.d.). While humans possess capabilities hard payloads do not, they bring their own set of limitations.

One of the commonly accepted characteristics of life is the production of waste; during a long-term mission such as the journey to Mars and back, biological waste will be an influential factor. Approximately 128 g of feces wet mass and 1.4 L (59 g dry solids) of urine are produced by each human daily (Rose, Parker, Jefferson, & Cartmell, 2015). With a crew of six and an estimated Earth-Mars transfer of 202 days (Horneck et al., 2006), such approximations predict 155 kg of feces wet mass and 1,697 L of urine (72 kg dry solids) produced upon Martian arrival. Consequently, biological waste will be a factor that cannot be ignored. Given that waste and radiation are two immovable constants, this paper aims to explore the idea of using human waste as a radiation shield.

While urine is primarily recycled for potable liquid and oxygen production (Jr., Carter, & Higbie, n.d.), feces have not been similarly utilized. For expeditions outside Low Earth Orbit (LEO), storm shelters will likely integrate compacted foodstuffs, equipment and biowaste (Simon et al., n.d.). Therefore, this investigation will explore the comparative qualities of urine and feces as elements of such a shield. Urine is primarily composed of water, which will likely result in no significant difference in attenuation compared to pure water. Human feces are composed of approximately 75% water and 25% solid matter (Britannica, 2002), which suggests a possibly significant negative effect on its attenuation properties after desiccation. Because the high energies required to directly simulate SPE particles would be impractical to achieve in the Space Environment Laboratory of California Polytechnic State University, San Luis Obispo, a surrogate must take the place of a particle accelerator.

While the β decay (two protons and two neutrons) of radioisotopes (RIs) occurs at orders of magnitude below that of SPEs, they are composed of similar particles: protons (Simon et al., n.d.); therefore, they may produce similar exposure effects on the leeward side of a shield. In order to determine if RIs may be used as SPE surrogates, they will be tested against shielding scenarios predicted by NASA's Online Tool for the Assessment of Radiation in Space (OLTARIS). The RIs that can replicate the trends to a statistically significant level will be used for the aforementioned experiments.

Materials & Methods

Hardware Specifications and Data Analysis

A glass vial of 0.87 in. inner diameter was placed in front of the Geiger-Mueller tube window (GMW) for all tests, which provided an effective shield thickness of the vial's inner diameter. The Geiger counter used was a Vernier® Digital Radiation Monitor connected to a Vernier® LabQuest Mini. Count measurements were collected at 10 s intervals for the respective timespans. Counts were then multiplied by a factor of six to convert to Counts per Minute (CPM). Statistical analysis was performed on the CPM data with MATLAB R2015b and Microsoft® Excel 2016.

Radioisotope Specifications

Three radioisotopes used included: 1 μCi . Co-60, 5 μCi . Cs-137 and 0.1 μCi . Po-210. All were in the form of discs acquired from Spectrum Technologies.

Radioisotope Surrogate Validation

The shielding scenario used for RI surrogate verification was a comparison of Aluminum-High Density Polyethylene (Al-HDPE) and High Density Polyethylene-Aluminum (Slaba et al., 2011). Two layers of 0.002 in. thick Al sheeting were combined with one layer of 0.004 in. thick HDPE to create a 0.008 in. thick Al-HDPE shield that was placed adjacent to the GMW on a Vernier Digital Radiation Monitor (**Fig. 1**). Each radioisotope (Co-60, Cs-137, and Po-210) was placed one in away and two 90 s exposures were taken for each shield configuration: Al-HDPE-GMW and HDPE-Al-GMW. A two-sample t-test at $\alpha = 0.05$ was run for each scenario pair to determine if any differences were significant. The RI(s) that were validated as surrogates were then used for the urine and feces tests.

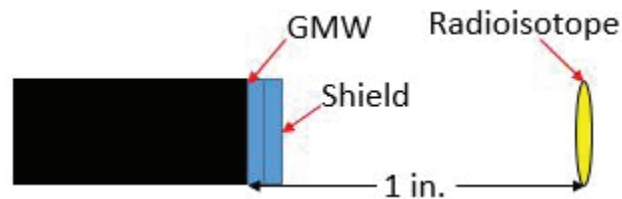


Figure 1: RI surrogate validation schematic. The 0.008 in. shield (both orientations) is adjacent to the GMW that is 1 in. from the RI (Co-60, Cs-137 and Po-210).

Urine-Water Shielding

The validated RI was placed three in. from the GMW and the vial of urine/water stood adjacent to the GMW. The vial was constant through all liquid tests, allowing analysis of CPM without compensation for the glass walls. The vial was filled with water and a 10 min exposure was recorded; the same procedure was followed for urine, which was used within 10 min of sample collection. Both fluids were above the top of the GMW, ensuring complete coverage relative to the RI (**Fig. 2**). A two-sample t-test at $\alpha = 0.05$ was run between both scenarios to determine if there were any significant differences in shielding effectiveness.

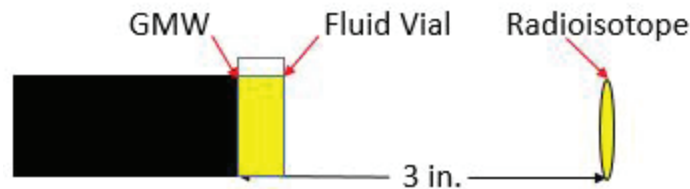


Figure 2: Urine/Water test schematic. The fluid-filled vial stood adjacent to the GMW, and the RI 3 in. from the GMW

Hydrated - Dehydrated Feces Shielding

The validated RI was placed three in. from the GMW and the vial of feces stood adjacent to the GMW. The vial was constant through both tests, allowing analysis of CPM without compensation for the glass walls. The vial was filled with feces and a 10 min exposure was recorded. The feces were left in the vial, placed in a vacuum chamber and exposed to pressures of approximately 490 mTorr. After desiccation, the sample was removed from the vacuum chamber and a 10 min exposure was recorded. Both feces levels were above the top of the GMW, ensuring complete coverage relative to the RI (**Fig. 3**). A two-sample t-test at $\alpha = 0.05$ was run between both scenarios to determine if there were any significant differences in shielding effectiveness.

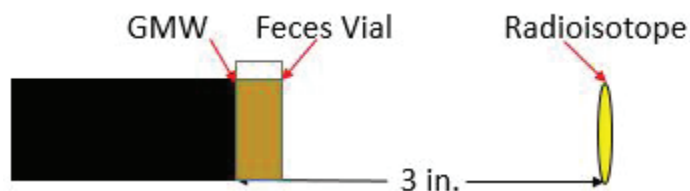


Figure 3: Hydrated and dehydrated feces test schematic. The feces-filled vial stood adjacent to the GMW, and the RI 3 in. from the GMW.

Results

Radioisotope Surrogate Validation

Out of the three RIs tested (Co-60, Cs-137 and Po-210), only Cs-137 exhibited a difference between Al-HDPE and HDPE-Al shields ($p < 0.05$). As shown in **Fig. 4**, the Al-HDPE orientation reduced CPM on the leeward side of the shield compared to the HDPE-Al orientation. Cs-137, being the only β emitter that mirrored the OLTARIS results (Slaba et al., 2011), positioned it as the RI surrogate for further tests.

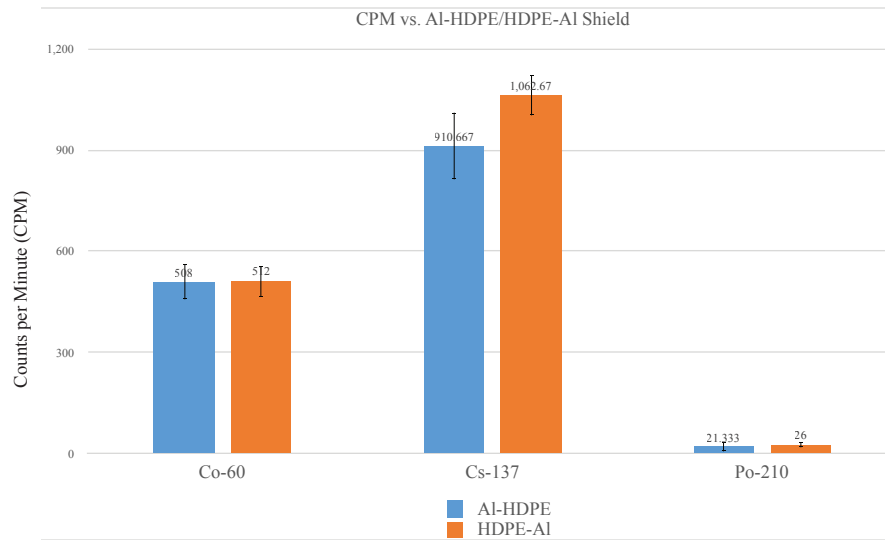


Figure 4: CPM (\pm S.E.) comparison between Al-HDPE and HDPE-Al shield orientations for Co-60, Cs-137 and Po-210. Only Cs-137 produced a difference ($p < 0.05$) in shield effectiveness based on orientation.

Urine-Water Shielding

No difference was found between urine and water shields of the same thickness (Fig. 5 - $p < 0.05$).

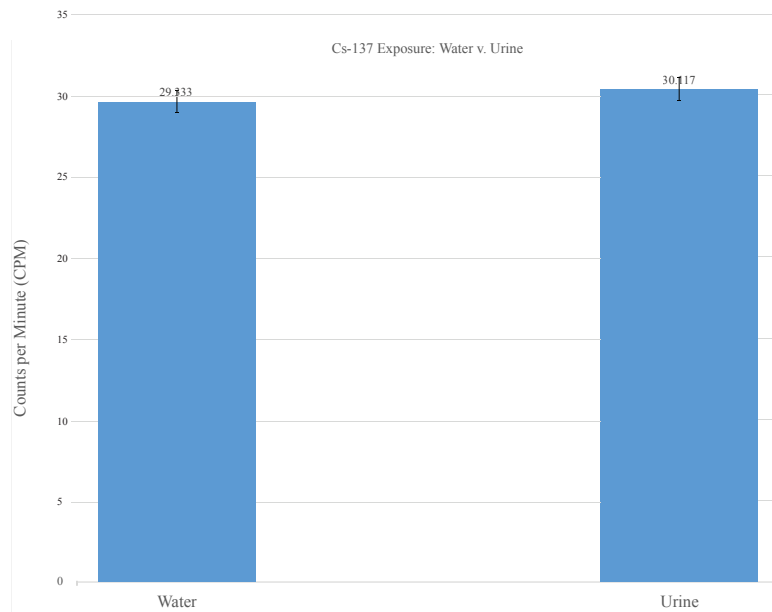


Figure 5: CPM (\pm S.E.) comparison between water and urine shields ($t = 0.87$ in.). No difference ($p < 0.05$) found between CPM on leeward side of shields.

Hydrated-Dehydrated Feces Shielding

No difference was found between hydrated and dehydrated feces shields of the same thickness (Fig. 6 - $p < 0.05$).

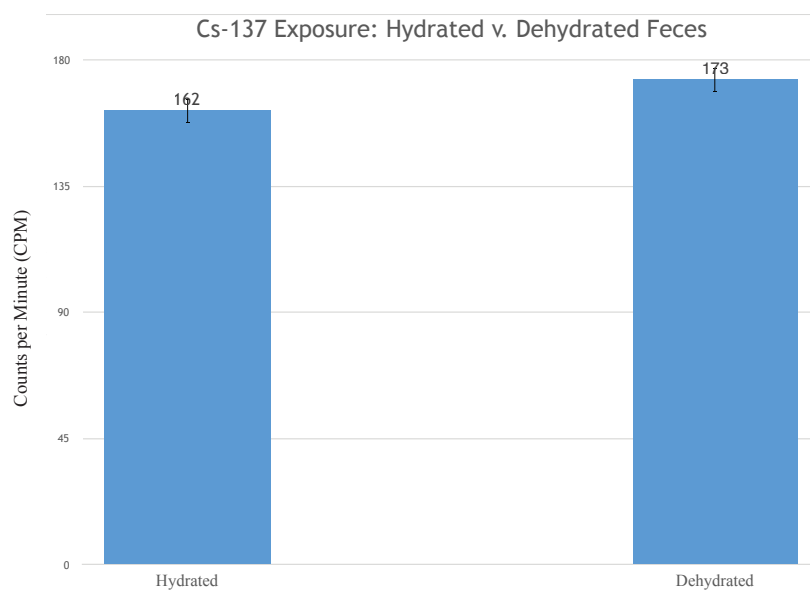


Figure 6: CPM (\pm S.E.) comparison between hydrated and dehydrated feces shields ($t = 0.87$ in.). No difference ($p < 0.05$) found between CPM on leeward side of shields.

There was a 33% decrease in mass from hydrated to dehydrated feces (**Table 1**).

| Hydrated (g) | Dehydrated (g) | Difference (%) |
|--------------|----------------|----------------|
| 14.6 | 9.75 | 33 |

Table 1: Mass of hydrated and dehydrated feces used for shielding

Discussion

This work encompassed the validation of a radioisotope surrogate to simulate SPE particles through comparative testing of radiation attenuation amongst both urine vs. water and hydrated vs. dehydrated feces shields.

The data mirroring behavior predicted in OLTARIS (Slaba et al., 2011) supports the use of Cs-137 as a stand-in for SPE particles during the subsequent shielding experiments. The lack of difference in shielding effectiveness found between urine and water suggests that urine carries more value in water reclamation operations than in radiation shielding. Given the corrosive nature of brine produced during water reclamation (Jr. et al., n.d.), it is unlikely that the solid components of urine would yield a worthwhile return in the form of radiation shielding.

The lack of difference in shielding effectiveness between hydrated and dehydrated feces suggests that the hydration level of feces should not be considered a significant factor in radiation shielding operations. This result may influence decisions encompassed by the holistic nature of space travel: the moisture contained in feces could be reclaimed and the remaining solid mass integrated into radiation shields, without fear of detrimental influences on shielding effectiveness.

Conclusion

There was no difference in radiation attenuation effectiveness between equal thickness shields of urine vs. water and hydrated vs. dehydrated feces. These results suggest that the reclamation of water from human biowaste does not detract from the waste's potential as a radiation shield.

While Cs-137 was validated as a small scale SPE surrogate, the conclusions reached in this paper may be strengthened by experiments involving higher power sources with greater particle specificity, such as HIMAC (Marc M. Cohen, n.d.).

References

Encyclopedia Britannica. (2002). Feces.

Horneck, G., Facius, R., Reichert, M., Rettberg, P., Seboldt, W., Manzey, D., Gerzer, R. (2006). Humex, a study on the survivability and adaptation of humans to long-duration exploratory missions, part ii: Missions to mars. *Advances in Space Research*, 38(4), 752 -759. Retrieved from <http://www.sciencedirect.com/science/article/pii/S0273117705008653> (Mercury, Mars and Saturn) doi: <http://dx.doi.org/10.1016/j.asr.2005.06.072>

Jr., D. E. L., Carter, D. L., & Higbie, S. (n.d.). Development of an advanced recycle filter tank assembly for the iss urine processor assembly.

Langford, M. (n.d.). Why is space radiation an important concern for human spaceflight? Retrieved from <https://srag.jsc.nasa.gov/spaceradiation/why/why.cfmz>

Miller, J., Cohen, M. M., Parodi, J. (n.d.). Water walls radiation shielding: Preliminary beam testing of ersatz solid waste simulant.

Rose, C., Parker, A., Jefferson, B., & Cartmell, E. (2015, Sep). The characterization of feces and urine: A review of the literature to inform advanced treatment technology. *Critical Reviews in Environmental Science Technology*, 45(17), 1827–1879.

Simon, M. A., Cloudsley, M., & Walker, S. (n.d.). Habitat design considerations for implementing solar particle event radiation protection.

Slaba, T. C., McMullen, A. M., Thibeault, S. A., Sandridge, C. A., Cloudsley, M. S., & Blattnig, S. R. (2011). Oltaris: An efficient web-based tool for analyzing materials exposed to space radiation.

About the Author

Noah Falck is a Mustang alumnus who graduated in Winter 2017 with a B.S. in Aerospace Engineering (Astronautics) and minors in Biology and Military Science. If he is not working on the Jeep or cooking, he can be found out running trails and spending time with our equine friends. He believes that exploration is at the core of what it means to be a human and aims to study Bioastronautics after service in the Army.

Acknowledgements

I would like to thank Dr. Abercromby for providing vacuum chamber facilities at the Space Environments Laboratory and the Aerospace Engineering Department for funding to acquire the radioisotopes and data collection hardware. Additionally, Dr. Howes of the Biological Sciences Department aided in hardware supply and helped train me in the process of experimental design and biological best practices.

III. Special Feature

A Letter from Dean Phil Bailey

The announcement of a \$110 million pledge from Bill and Linda Frost on May 3, 2017 changed everything for Cal Poly's College of Science and Mathematics. We are now able to pursue a bold vision of establishing the College as outstanding and nationally recognized in academic and student research programs. And we have the resources to do so thanks to Bill and Linda Frost.

Both Bill and Linda are California State University graduates. Linda Frost graduated from San Jose State University with a degree in biology and Bill Frost came to Cal Poly as a transfer student in the winter of 1970, earning his degree in biochemistry. I arrived at Cal Poly as a 26-year-old assistant professor in the fall of 1969, not long before Bill. We have been in contact ever since and worked together to define this historic gift, the largest in the history of the California State University system and one of the largest in the history of public universities in California.

Bill had an excellent undergraduate research experience at Cal Poly and credits it with giving him the skills and confidence "to define and solve problems as well as actually develop and implement solutions." These attributes served him well as he successfully formed his own businesses. Bill and Linda want to give current and future students that same experience of discovery that Bill had as a Cal Poly student.

Engaging in real-world research with faculty mentors presents students with questions yet to be answered. Finding the solutions engages one's curiosity, imagination, and creative and critical thinking skills. Interdisciplinary approaches and entrepreneurship can be involved as well as both independent thinking and group collaboration. Participating in research is a special educational experience for undergraduate students.

For well over a decade, Bill supported four students in research each summer. Beginning in 2014, he dramatically increased his support (which we matched with College Based Fee) and, since then, we have had approximately 200 students performing research with their faculty mentors full time each summer in the College of Science and Mathematics. This is just the beginning. Because of the Frosts' thoughtful generosity to the College of Science and Mathematics:

- Imagine that we are now able to attract, retain, and support students with generous stipends and scholarships to pursue degrees in the College of Science and Mathematics and engage in undergraduate research.
- Imagine that we can now provide our faculty with release time to more fully mentor our students in research and involve them in presentations and professional meetings and as co-authors on peer-reviewed publications.
- Imagine that we will now be able to continuously ensure that we have the best equipment and instrumentation needed for our students to perform real-life research.
- And, imagine that, in just a few years, we will have a new building with over 18,000 square feet of additional facilities for student-faculty research.

The College of Science and Mathematics has changed forever.

Phil Bailey, Dean

College of Science and Mathematics



SYMPOSIUM
Student Journal of Science & Math



**Thank you to all the individuals who have developed, contributed, supported, and been a part of *Symposium*.
You made this dream a reality.**

