THE EXPECTED VARIATION OF RANDOM BOUNDED INTEGER SEQUENCES OF FINITE LENGTH

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From the enumerative generating function of an abstract adjacency statistic, we deduce the mean and variance of the variation on random permutations, rearrangements, compositions, and bounded integer sequences of finite length.

1. Introduction

When the finite sequence of integers \( w = 1, 3, 2, 2, 4, 3 \) is sketched as below,

\[
  w = \begin{array}{c}
    1 \\
    2 \\
    3 \\
    4 \\
    3 \\
    \end{array}
\]

its most compelling aspect is its vertical variation, that is, the sum of the vertical distances between its adjacent terms. Denoted by \( \text{var} w \), the vertical variation of the sequence in (1.1) is \( \text{var} w = 2 + 1 + 0 + 2 + 1 = 6 \). Our purpose here is to compute the mean and variance of \( \text{var} \) on four classical sets of combinatorial sequences.

To formalize matters and place our problem in the context of other work, let \([m]^n\) denote the set of sequences \( w = x_1x_2 \cdots x_n \) of length \( n \) with each \( x_i \in \{1, 2, \ldots, m\} \). For a real-valued function \( f \) on \([m]^2\), the \( f \)-adjacency number of \( w = x_1x_2 \cdots x_n \in [m]^n \) is defined to be

\[
  \text{adf} w = \sum_{k=1}^{n-1} f(x_kx_{k+1}).
\]
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Table 1.1

<table>
<thead>
<tr>
<th>Sequences</th>
<th>Expected value of var</th>
<th>Variance of var</th>
</tr>
</thead>
<tbody>
<tr>
<td>$S_n$</td>
<td>$n^2 - 1$</td>
<td>$(n - 2)(n + 1)(4n - 7)$</td>
</tr>
<tr>
<td>$[m]^n$</td>
<td>$(n - 1)(m^2 - 1)$</td>
<td>$(m^2 - 1)(6m^2n + 6n - 7m^2 - 2)$</td>
</tr>
<tr>
<td>$R_n(i)$</td>
<td>$\frac{2}{n} \sum_{1 \leq i &lt; j \leq m} (y - x)_{i,j}$</td>
<td>See (3.10)</td>
</tr>
<tr>
<td>$C_n(m)$</td>
<td>$\frac{2(n - 1)}{(m - 1)^2} \sum_{x=1}^{\lfloor m/2 \rfloor} (m - 2x)^{2-\frac{1}{2}}$</td>
<td>See (3.18)</td>
</tr>
</tbody>
</table>

Some specializations of the $f$-adjacency number have been considered elsewhere. For instance, if $f(xy)$ is 1 when $x < y$ and 0 otherwise, then adf $w$ is known as the rise number of $w$ [1, 3, 4]. For the selection $f(xy) = |y - x|$, adf $w$ = var $w$. In a sorting problem of computer science, Levopoulos and Petersson [5] introduced the related notion of oscillation (var $w - n + 1$) as a measure of the presortedness of a sequence of $n$ distinct numbers. In [6], compositions were enumerated by their ascent variation, the $f$-adjacency statistic induced by $f(xy) = y - x$ if $x < y$ and 0 otherwise. For the case $f(xy) = h(|y - x|)$ where $h$ is a linear, convex, or concave increasing real-valued function, Chao and Liang [2] described the arrangements of $n$ distinct integers for which adf achieves its extreme values.

Besides considering the distribution of var on the set $[m]^n$, we also consider it on the set of rearrangements $R_n(i_1, i_2, \ldots, i_m)$ consisting of sequences of length $n = i_1 + i_2 + \cdots + i_m$ which contain $i_l$ exactly $i_l$ times, on the set of permutations $S_n = R_n(1, 1, \ldots, 1)$ of $\{1, 2, \ldots, n\}$, and on the set of compositions of $m$ into $n$ parts $C_n(m) = \{x_1, x_2, \ldots, x_n \in [m]^n : x_1 + x_2 + \cdots + x_n = m\}$. For $m,n \geq 2$, Table 1.1 displays the mean and variance of var on these four sets. The $k$th falling factorial of $n$ is $n^{\underline{k}} = n(n - 1) \cdots (n - k + 1)$, $i = (i_1, i_2, \ldots, i_m)$, and, for $r$ a real number, $[r]$ denotes the greatest integer less than or equal to $r$. The results in Table 1.1 are new. David and Barton [3, Chapter 10] present the distributions of several statistics (some $f$-adjacency numbers, some not) primarily on permutations. We also note that Tiefenbruck [7] derived a generating function for compositions with bounded parts by a close relative of var. We leave open questions concerning the asymptotic behavior of var.

2. Enumerative factorial moments for $f$-adjacencies

Before working specifically with var, we discuss the enumerative generating function for adf on sequences as developed by Fédu and Rawlings [4]. Let $[m]^*$ denote the set of sequences of $1, 2, \ldots, m$ of finite length (including the empty sequence of length 0). For $w = x_1 x_2 \cdots x_n \in [m]^*$, we define $Z^w = z_{x_1} z_{x_2} \cdots z_{x_n}$. The enumerative generating function for adf over $[m]^*$ is then defined to be $G(p) = \sum_{w \in [m]^*} p^{\text{adf}^w} Z^w$.

By manipulating $G(p)$, we will obtain all of the information in Table 1.1 (and more). As a brief outline of our approach, note that the coefficient of $p^k z_1^{i_1} z_2^{i_2} \cdots z_m^{i_m}$ in $G(p)$ is
just the number of rearrangements \( w \) in \( R_n(i) \) with \( \text{adf} \ w = k \). Thus, by dividing the coefficient of \( z_1^i z_2^j \cdots z_m^k \) in \( G' \) by the cardinality of \( R_n(i) \), we will obtain the mean of \( \text{adf} \). So, in general, we compute the \( d \)th enumerative factorial moment \( G^{(d)}(1) = \sum_{w \in [m]^{n}} (\text{adf} \ w)^{d} \ Z^{w} \).

From the work of Fédou and Rawlings [4], it follows that

\[
G(p) = \frac{1}{D(p)},
\]

where

\[
D(p) = 1 - \sum_{n \geq 1} \sum_{x_1, \ldots, x_n \in [m]^n} Z^{x_1 \cdots x_n} \prod_{k=1}^{n-1} (p f(x_k, x_{k+1}) - 1).
\]

Examples are presented in [4, 6] for which \( D \) has a closed form. We do not know a closed form for \( D \) when \( \text{adf} = \text{var} \) (that is, when \( f(x, y) = |y - x| \)). Nevertheless, (2.1) is still useful in computing the mean and variance of \( \text{var} \).

Although the formula for taking the \( d \)-fold derivative with respect to \( p \) of a function of the form in (2.1) is known, we provide a short derivation. To avoid the quotient and chain rules, rewrite (2.1) as \( GD = 1 \). Differentiating the latter \( d \) times, \( d \geq 1 \), and dividing by \( d! \) gives

\[
\sum_{j=0}^{d} \frac{G^{(d-j)} D^{(j)}}{(d-j)! j!} = 0.
\]

To solve for \( G^{(d)} \), consider the system

\[
\frac{G^{(d)}}{d!} \frac{D^{(0)}}{0!} + \frac{G^{(d-1)}}{(d-1)!} \frac{D^{(1)}}{1!} + \frac{G^{(d-2)}}{(d-2)!} \frac{D^{(2)}}{2!} + \cdots + \frac{G^{(0)}}{0!} \frac{D^{(d)}}{d!} = 0,
\]

\[
\frac{G^{(d-1)}}{(d-1)!} \frac{D^{(0)}}{0!} + \frac{G^{(d-2)}}{(d-2)!} \frac{D^{(1)}}{1!} + \cdots + \frac{G^{(0)}}{0!} \frac{D^{(d-1)}}{(d-1)!} = 0,
\]

\[
\vdots
\]

\[
\frac{G^{(1)}}{1!} \frac{D^{(0)}}{0!} + \frac{G^{(0)}}{0!} \frac{D^{(1)}}{1!} = 0,
\]

\[
\frac{G^{(0)}}{0!} \frac{D^{(0)}}{0!} = 1,
\]

where the top \( d \) equations arise from repeated application of (2.3). Cramer’s rule applied
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to the above system yields

\[
G^{(d)} = \frac{(-1)^d}{d!} \begin{pmatrix}
\frac{D^{(1)}}{1!} & \frac{D^{(2)}}{2!} & \frac{D^{(3)}}{3!} & \cdots & \frac{D^{(d)}}{d!} \\
0 & \frac{D^{(0)}}{1!} & \frac{D^{(1)}}{2!} & \cdots & \frac{D^{(d-1)}}{(d-1)!} \\
\vdots & \vdots & \vdots & \ddots & \vdots \\
0 & \cdots & 0 & \frac{D^{(0)}}{0!} & \frac{D^{(1)}}{1!}
\end{pmatrix}
\]  

(2.5)

which, when expanded, implies that

\[
G^{(d)} = \sum_{\nu=1}^{d} \left( \frac{(-1)^\nu}{D^{\nu+1}} \right) \sum_{j_1+\cdots+j_\nu=d, j_k \geq 1} \binom{d}{j_1 \cdots j_\nu} D^{(j_1)} \cdots D^{(j_\nu)}. 
\]  

(2.6)

To determine the enumerative factorial moment \(G^{(d)}(1)\), we see from (2.2) that

\[
D^{(j)}(1) = -\sum_{r=2}^{j+1} D^{(j)}_r, 
\]

(2.7)

where

\[
D^{(j)}_r = \sum_{x_1, \ldots, x_r \in [m]^j} \sum_{l_1, \ldots, l_{r-1} = j_k \geq 1} \binom{j}{l_1 \cdots l_{r-1}} \prod_{k=1}^{r-1} f(x_k x_{k+1})^j. 
\]  

(2.8)

For instance,

\[
D'_2 = \sum_{x, y \in [m]^2} f(xy)zxzy, \quad D'_2 = \sum_{x, y \in [m]^2} f(xy)^2z_xz_y, \\
D'_3 = 2 \sum_{x, y, z \in [m]^3} f(vx)f(xy)z_xz_yz_x. 
\]  

(2.9)

Further setting \(\tilde{j} = (j_1, \ldots, j_\nu)\), \(s(\tilde{j}) = j_1 + \cdots + j_\nu\),

\[
\binom{d}{\tilde{j}} = \binom{d}{j_1 \cdots j_\nu}, \quad \text{and} \quad D^{(j_1)}_{r_1} \cdots D^{(j_\nu)}_{r_\nu} = \sum_{r_1, \ldots, r_\nu = \mu}^{r_1 + \cdots + r_\nu = \mu} D^{(j_1)}_{r_1} \cdots D^{(j_\nu)}_{r_\nu}, 
\]

(2.10)

it follows from (2.6) and (2.7) that

\[
G^{(d)}(1) = \sum_{\nu=1}^{d} \frac{1}{D^{\nu+1}(1)} \sum_{s(\tilde{j})=d, j_k \geq 1} \binom{d}{\tilde{j}} \sum_{\mu=2\nu}^{d+\nu} D^{(j_1)}_{r'_1} \cdots D^{(j_\nu)}_{r'_\nu}. 
\]  

(2.11)
As \( D(1) = 1 - (z_1 + \cdots + z_m) \), extracting the contributions made by all \( w \in [m]^n \) from both sides of (2.11) gives the \( d \)th enumerative factorial moment of \( \text{adf} \) over \( [m]^n \) as

\[
\sum_{w \in [m]^n} (\text{adf} \ w)^d Z^w = \sum_{y=1}^{d} \sum_{\substack{s(j)=d \\atop j \geq 1}} \binom{d}{y} \sum_{\mu=2y}^{n+y-\mu} \left( \sum_{i=1}^{m} z_i \right)^{n-\mu} D^{(j)}_p \tag{2.12}
\]

valid for \( d \geq 1 \). When \( d = 1,2 \), (2.9) and (2.12) imply that

\[
\sum_{w \in [m]^n} \text{adf} \ w Z^w = (n-1) \left( \sum_{i=1}^{m} z_i \right)^{n-2} \sum_{xy \in [m]^2} f(xy) z_x z_y \tag{2.13}
\]

and that

\[
\sum_{w \in [m]^n} (\text{adf} \ w)^2 Z^w = (n-1) \left( \sum_{i=1}^{m} z_i \right)^{n-2} \sum_{xy \in [m]^2} (f(xy))^2 z_x z_y
+ 2(n-2) \left( \sum_{i=1}^{m} z_i \right)^{n-3} \sum_{vxy \in [m]^3} f(vx) f(xy) z_v z_x z_y \tag{2.14}
\]

\[+ (n-2)(n-3) \left( \sum_{i=1}^{m} z_i \right)^{n-4} \left( \sum_{xy \in [m]^2} f(xy) z_x z_y \right)^2. \]

3. Discussion of Table 1.1

The entries in Table 1.1 are consequences of (2.13) and (2.14) with \( f(xy) = |y-x| \) and with appropriate substitutions for \( Z \). For the mean and variance of \( \text{var} \) on the set of bounded sequences \( [m]^n \), put \( z_i = 1 \) for \( 1 \leq i \leq m \). Noting that

\[
\sum_{xy \in [m]^2} |y-x| = \sum_{1 \leq x < y \leq m} 2(y-x) = 2 \binom{m+1}{3}, \tag{3.1}
\]

it follows from (2.13) that the mean of \( \text{var} \) on \( [m]^n \) is

\[
\frac{1}{m^n} \sum_{w \in [m]^n} \text{var} \ w = \frac{2(n-1)m^{n-2}}{m^n} \binom{m+1}{3} = \frac{(n-1)(m^2-1)}{3m}. \tag{3.2}
\]

As

\[
\sum_{xy \in [m]^2} |y-x|^2 = 4 \binom{m+1}{4} \tag{3.3}
\]
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and as

\[
\sum_{v < x < y \leq m} |x - v||y - x| = \sum_{1 \leq v < x < y \leq m} 2(x - v)(y - x) \\
+ \sum_{1 \leq x < y \leq m} 4(v - x)(y - x) - \sum_{1 \leq x < y \leq m} 2(y - x)^2
\]

\[
= \frac{7m^2 - 8}{10} \binom{m + 1}{3},
\]

(3.4)

(2.14) implies that

\[
\frac{1}{m^n} \sum_{w \in [m]^n} (\text{var } w)^2 = \frac{4(n - 1)}{m^2} \left( \frac{m + 1}{4} \right) + \frac{(n - 2)(7m^2 - 8)}{5m^3} \left( \frac{m + 1}{3} \right) \]
\[
+ \frac{4(n - 2)(n - 3)}{m^4} \left( \frac{m + 1}{3} \right)^2.
\]

(3.5)

Then, subbing the last result into

\[
\frac{1}{m^n} \sum_{w \in [m]^n} (\text{var } w)^2 + \frac{(n - 1)(m^2 - 1)}{3m} - \left( \frac{(n - 1)(m^2 - 1)}{3m} \right)^2
\]

(3.6)

and simplifying gives the variance of var as recorded in Table 1.1.

For \(R_n(\vec{i})\), extracting the coefficient of \(z_{i_1}^{i_1} z_{i_2}^{i_2} \cdots z_{i_m}^{i_m}\) from (2.13) leads to

\[
\sum_{w \in R_n(\vec{i})} \text{var } w = 2(n - 1) \sum_{1 \leq i_x < i_y \leq m} (y - x) \left( \binom{n}{i_1, i_2 \cdots i_m} \right).
\]

(3.7)

As the cardinality of \(R_n(\vec{i})\) is

\[
\binom{n}{i_1 i_2 \cdots i_m} = \binom{n}{\vec{i}},
\]

(3.8)

it follows that the mean of var on \(R_n(\vec{i})\) is

\[
\binom{n}{\vec{i}}^{-1} \sum_{w \in R_n(\vec{i})} \text{var } w = \frac{2}{n} \sum_{1 \leq i_x < i_y \leq m} (y - x) i_x i_y.
\]

(3.9)

Let \(\vec{i}_r = (i_1, \ldots, i_r - 1, \ldots, i_n)\). For example, \((3,2,1,4)_{1}^{3} = (3,1,-1,4)\). The variance on \(R_n(\vec{i})\) is then

\[
\binom{n}{\vec{i}}^{-1} \sum_{w \in R_n(\vec{i})} \text{var } w = \frac{2}{n} \sum_{1 \leq i_x < i_y \leq m} (y - x) i_x i_y - \left( \frac{2}{n} \sum_{1 \leq i_x < i_y \leq m} (y - x) i_x i_y \right)^2,
\]

(3.10)
where, upon extraction of the coefficient of $z_1^i z_2^j \ldots z_m^w$ from (2.14), we have

$$
\sum_{w \in R_n(i)} (\text{var } w)^2 = (n - 1) \sum_{1 \leq x, y \leq m} |y - x|^2 \left( \frac{n - 2}{i_{\{x \backslash y\}}^2} \right) + 2(n - 2) \sum_{1 \leq v, x, y \leq m} |x - v||y - x| \left( \frac{n - 3}{i_{\{v \backslash x\}}^2} \right) \nonumber
\sum_{1 \leq u, v, x, y \leq m} |v - u||y - x| \left( \frac{n - 4}{i_{\{u \backslash v \backslash x\}}^2} \right).
$$

(3.11)

The permutation entries in Table 1.1 follow from (3.9) and (3.10). Selecting $m = n$ and $i_k = 1$ for $1 \leq k \leq n$ in (3.9) reveals the mean of var on $S_n$ as

$$
\frac{1}{n!} \sum_{w \in S_n} \text{var } w = \frac{2}{n} \sum_{1 \leq x < y \leq n} (y - x) = \frac{2}{n} \left( \frac{n+1}{3} \right) = \frac{n^2 - 1}{3}.
$$

(3.12)

From (3.11), with $m = n$ and $i_k = 1$ for $1 \leq k \leq n$,

$$
\sum_{w \in S_n} (\text{var } w)^2 = (n - 1)! \sum_{1 \leq x, y \leq n} |y - x|^2 + 2(n - 2)! \sum_{1 \leq v, x, y \leq n} |x - v||y - x| \nonumber
\sum_{1 \leq u, v, x, y \leq m} |v - u||y - x| \left( \frac{n - 4}{15} \right) (n - 2)! (10n^2 + 14n - 27) \left( \frac{n + 1}{4} \right).
$$

(3.13)

So the variance of var on $S_n$ is

$$
\frac{1}{n!} \sum_{w \in S_n} \text{var } w^2 + \frac{n^2 - 1}{3} - \left( \frac{n^2 - 1}{3} \right)^2 = \frac{(n - 2)(n + 1)(4n - 7)}{90}.
$$

(3.14)

For $w = x_1 \ldots x_n \in [m]^n$, let $\|w\| = x_1 + \cdots + x_n$. For the composition results in Table 1.1, set $z_k = q^k$ for $1 \leq k \leq m$. Then (2.13) implies that

$$
\sum_{w \in [m]^n} \text{var } w \|w\| = (n - 1)q^{n-2} \left( \frac{1 - q^m}{1 - q} \right) \sum_{1 \leq x, y \leq m} |y - x| q^{x+y}.
$$

(3.15)
and (2.14) leads to

\[
\sum_{w \in [m]^n} (\text{var } w)^2 \|w\| = (n-1)q^{n-2}\left(1 - \frac{q^m}{1 - q}\right)^{n-2} \sum_{1 \leq x, y \leq m} |y - x|^2 q^{x+y} \\
+ 2(n-2)q^{n-3}\left(1 - \frac{q^m}{1 - q}\right)^{n-3} \sum_{1 \leq v, x, y \leq m} |x - v||y - x| q^{v+x+y} \\
+ (n-2)(n-3)q^{n-4}\left(1 - \frac{q^m}{1 - q}\right)^{n-4} \sum_{1 \leq u, v, x, y \leq m} |v - u||y - x| q^{u+v+x+y}.
\]

(3.16)

Extracting the coefficient of \(q^m\) from (3.15) to obtain

\[
\sum_{w \in C_n(m)} \text{var } w = 2(n-1) \sum_{1 \leq x < y \leq m} (y - x) \binom{m-1-x-y}{n-3} \\
= 2(n-1) \sum_{1 \leq x \leq \lfloor m/2 \rfloor} \binom{m-2x}{n-1} 
\]

(3.17)

and then dividing by the cardinality \(n^{-1}\) of \(C_n(m)\) gives the mean of var as stated in Table 1.1. The variance is

\[
\left( \frac{m-1}{n-1} \right)^{-1} \sum_{w \in C_n(m)} \text{var } w^2 + \frac{2(n-1)}{(m-1)^{n-1}} \sum_{1 \leq x \leq \lfloor m/2 \rfloor} (m-2x)^{n-1} \\
- \left( \frac{2(n-1)}{(m-1)^{n-1}} \sum_{1 \leq x \leq \lfloor m/2 \rfloor} (m-2x)^{n-1} \right)^2, 
\]

(3.18)

where, pulling the coefficient of \(q^m\) from (3.16), we have

\[
\sum_{w \in C_n(m)} (\text{var } w)^2 = (n-1) \sum_{1 \leq x, y \leq m} |y - x|^2 \binom{m-1-x-y}{n-3} \\
+ 2(n-2) \sum_{1 \leq v, x, y \leq m} |x - v||y - x| \binom{m-1-v-x-y}{n-4} \\
+ (n-2)(n-3) \sum_{1 \leq u, v, x, y \leq m} |v - u||y - x| \binom{m-1-u-v-x-y}{n-5}.
\]

(3.19)

The sums in (3.19) are marginally simplified. For instance,

\[
\sum_{1 \leq x, y \leq m} |y - x|^2 \binom{m-1-x-y}{n-3} = 4 \sum_{1 \leq x \leq \lfloor m/2 \rfloor} \binom{m-2x}{n}.
\]

(3.20)
As a part of the second sum on the right-hand side of (3.19), we note that
\[
\sum_{1 \leq v < x < y \leq m} (x - v)(y - x) \left( \frac{m - 1 - v - x - y}{n - 4} \right)
\]
\[= \sum_{2 \leq x \leq \lfloor (m + 1)/2 \rfloor} \left( \frac{m - 3x + 1}{n} \right) - \left( \frac{m - 2x + 1}{n} \right) + x \left( \frac{m - 2x}{n - 1} \right). \tag{3.21}
\]
The four-fold sums arising in the last sum in (3.19) reduce to double sums.

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